

### TD 3: Security Assumptions (corrected version)

**Exercise 1.***Advantage(s)*

We consider two distributions  $D_0$  and  $D_1$  over  $\{0, 1\}^n$  and the following experiment.

$\mathcal{C}$	$\mathcal{A}$
sample $b \leftarrow U(0, 1)$ sample $x \leftarrow D_b$ send $x$ to $\mathcal{A}$	compute a bit $b'$ send $b'$ to $\mathcal{C}$
If $b = b'$ , say “Win”, else say “Lose”.	

We say that a PPT (Probabilistic, Polynomial-Time) algorithm  $\mathcal{A}$  is a *distinguisher* if there exists a non-negligible  $\varepsilon$  such that, in this experiment,  $\Pr[\text{Win}] \geq 1/2 + \varepsilon$ . The distributions  $D_0$  and  $D_1$  are said to be *indistinguishable* if there is no such distinguisher.

1. Show that this definition of indistinguishability is equivalent to the one seen during the lecture.

We will answer by giving two reductions.

Let  $\mathcal{A}$  be an attacker regarding the TD definition; we then consider  $\mathcal{A}'$ , defined exactly as  $\mathcal{A}$  except that it outputs  $b'$  rather than sending it to  $\mathcal{C}$ . We have

$$\begin{aligned}
 \text{Adv}(\mathcal{A}') &= |\Pr[\mathcal{A}' \xrightarrow{\text{Exp}_0} 1] - \Pr[\mathcal{A}' \xrightarrow{\text{Exp}_1} 1]| \\
 &= |\Pr[b' = 1|b = 0] - \Pr[b' = 1|b = 1]| \\
 &= |1 - \Pr[b' = 0|b = 0] - \Pr[b' = 1|b = 1]| \\
 &= |1 - 2\Pr[b' = 0 \text{ and } b = 0] - 2\Pr[b' = 1 \text{ and } b = 1]| \\
 &= |1 - 2\Pr(\text{Win})| \geq 2\varepsilon.
 \end{aligned}$$

Conversely, let  $\mathcal{A}'$  be an attacker regarding the definition seen in class, and define  $\mathcal{A}$  exactly as  $\mathcal{A}'$  except that, instead of outputting  $b'$ , it sends it to  $\mathcal{C}$ . Then, we have

$$\begin{aligned}
 \Pr(\text{Win}) &= \Pr[b' = 0 \text{ and } b = 0] + \Pr[b' = 1 \text{ and } b = 1] \\
 &= \frac{1}{2}(\Pr[b' = 0|b = 0] + \Pr[b' = 1|b = 1]) \\
 &= \frac{1}{2}(1 - \Pr[b' = 1|b = 0] + \Pr[b' = 1|b = 1]) \\
 &= \frac{1}{2}(1 + \text{Adv}(\mathcal{A}')).
 \end{aligned}$$

where we have assumed that  $\Pr[b' = 1|b = 1] \geq \Pr[b' = 1|b = 0]$ . Otherwise, we define  $\mathcal{A}$  exactly as  $\mathcal{A}'$  except that instead of outputting  $b'$  it sends  $1 - b'$  to  $\mathcal{C}$  (so that  $\Pr[b' = 1|b = 1] - \Pr[b' = 1|b = 0]$  is always positive and equal to  $\text{Adv}(\mathcal{A}')$ ).

*Remark.* For the reduction to work, we do not need to determine whether  $\Pr[\mathcal{A}' \xrightarrow{\text{Exp}_0} 1] - \Pr[\mathcal{A}' \xrightarrow{\text{Exp}_1} 1]$  is positive or negative. We just need to prove that whatever the case is, there exists an attacker  $\mathcal{A}$ .

2. A rebellious student decides to define a distinguisher as a PPT algorithm  $\mathcal{A}$  with  $\Pr[\text{Win}] \leq 1/2 - \varepsilon$  in the above experiment (rather than  $\geq 1/2 + \varepsilon$ ). Is this a revolutionary idea?

This is again equivalent to the previous definitions. Given an attacker  $\mathcal{A}$  for the game defined in this tutorial, we define  $\mathcal{A}'$  which acts exactly as  $\mathcal{A}$  except that it sends  $1 - b'$  to  $\mathcal{C}$  instead of  $b'$ . Then we have  $\Pr[\mathcal{A} \text{ win}] = 1 - \Pr[\mathcal{A}' \text{ win}]$ . And so  $\Pr[\mathcal{A} \text{ win}] \geq 1/2 + \varepsilon$  if and only if  $\Pr[\mathcal{A}' \text{ win}] \leq 1/2 - \varepsilon$ .

**Exercise 2.***Around the DDH assumption*

We recall the definition of the DDH assumption.

**Definition 1** (Decisional Diffie-Hellman distribution). Let  $\mathbb{G}$  be a cyclic group of (prime) order  $p$ , and let  $g$  be a public generator of  $\mathbb{G}$ . The decisional Diffie-Hellman distribution (DDH) is,  $D_{\text{DDH}} = (g^a, g^b, g^{ab}) \in \mathbb{G}^3$  with  $a, b$  sampled independently and uniformly in  $\mathbb{Z}/p\mathbb{Z} =: \mathbb{Z}_p$ .

**Definition 2** (Decisional Diffie-Hellman assumption). The decisional Diffie-Hellman assumption states that there exists no probabilistic polynomial-time distinguisher between  $D_{\text{DDH}}$  and  $(g^a, g^b, g^c)$  with  $a, b, c$  sampled independently and uniformly at random in  $\mathbb{Z}_p$ .

1. Does the DDH assumption hold in  $\mathbb{G} = (\mathbb{Z}_p, +)$  for  $p = \mathcal{O}(2^\lambda)$  prime?  $\text{☹}$  No. In this case, the  $D_{\text{DDH}}$  distribution is  $(a \cdot g, b \cdot g, (ab) \cdot g)$ . This can be distinguished from  $(ag, bg, cg)$  by computing the inverse of  $g$  (find a Bézout identity  $gu + pv = 1$  in logarithmic time), retrieving  $a, b$  and  $c$  and checking whether  $ab = c$  or not. This is always the case in the DDH, and the case with probability  $1/p$  in the uniform case. The advantage of a distinguisher returning the boolean value of  $ab = c$  is then  $\frac{p-1}{p}$ .
2. Same question for  $\mathbb{G} = (\mathbb{Z}_p^*, \times)$  of order  $p - 1$ , with  $p$  an odd prime.

$\text{☹}$  No, because  $p - 1$  (the order the group) is divisible by 2.

We know that  $x^{\frac{p-1}{2}} = 1$  if  $x \in \mathbb{Z}_p$  is a square and  $-1$  otherwise (it is actually the Legendre symbol:  $\left(\frac{x}{p}\right)$  and can be efficiently computed). So  $\left(\frac{g^a}{p}\right)$  gives us the parity of  $a$ , that is  $\left(\frac{g^a}{p}\right) = 1$  if  $a$  is even and  $\left(\frac{g^a}{p}\right) = -1$  if  $a$  is odd. Hence, if  $a$  is uniformly sampled in  $\{0, \dots, p-1\}$  (meaning that  $g^a$  is uniformly sampled in  $\mathbb{G}$ ), then  $\left(\frac{g^a}{p}\right)$  is uniformly distributed in  $\{-1, 1\}$ . But in the case of the DDH distribution, if  $a$  or  $b$  is even, then  $ab$  must be even too (or equivalently, if  $g^a$  or  $g^b$  is a square, then  $g^{ab}$  should be a square too). In the same way, if both  $a$  and  $b$  are odd, then  $ab$  must be odd.

This enables us to build the following distinguisher  $\mathcal{A}$ :

- Return DDH if  $\left(\frac{g^{ab}}{p}\right)$  is consistent with  $\left(\frac{g^a}{p}\right)$  and  $\left(\frac{g^b}{p}\right)$  (i.e.  $ab$  is odd and both  $a$  and  $b$  are odd, or  $ab$  is even and  $a$  or  $b$  is even);
- Return Unif otherwise.

Let us now compute the advantage of such a distinguisher.

$$\begin{aligned} \text{Adv}^{\text{DDH}}(\mathcal{A}) &= |\Pr[\mathcal{A} \rightarrow \text{DDH} \mid \text{DDH}] - \Pr[\mathcal{A} \rightarrow \text{DDH} \mid \text{Unif}]| \\ &= |1 - \Pr[\mathcal{A} \rightarrow \text{DDH} \mid \text{Unif}]| \end{aligned}$$

Our distinguisher returns Unif only if  $c$  is odd and either  $a$  or  $b$  is even or if  $c$  is even and both  $a$  and  $b$  are odd. But we have seen that these cases could not appear in the DDH distribution. So we have that  $\Pr[\mathcal{A} \rightarrow \text{DDH} \mid \text{DDH}] = 1$ .

It then remains to compute  $\Pr[\mathcal{A} \rightarrow \text{DDH} \mid \text{Unif}]$ . Given a Unif instance  $(g^a, g^b, g^c)$ , we have seen that  $\left(\frac{g^a}{p}\right)$ ,  $\left(\frac{g^b}{p}\right)$  and  $\left(\frac{g^c}{p}\right)$  are uniform in  $\{-1, 1\}$  because  $a, b, c$  are uniform in  $\{0, \dots, p-1\}$ . They are also independent because  $a, b$  and  $c$  are. So all eight possibilities for  $\left(\left(\frac{g^a}{p}\right), \left(\frac{g^b}{p}\right), \left(\frac{g^c}{p}\right)\right)$  have the same probability and we have

$$\begin{aligned} \Pr[\mathcal{A} \rightarrow \text{DDH} \mid \text{Unif}] &= \Pr\left[\left(\left(\frac{g^a}{p}\right), \left(\frac{g^b}{p}\right), \left(\frac{g^c}{p}\right)\right) = (1, 1, 1) \text{ or } (1, -1, 1) \text{ or } (-1, 1, 1) \text{ or } (-1, -1, -1)\right] \\ &= \frac{4}{8} = \frac{1}{2} \end{aligned}$$

To conclude, we have  $\text{Adv}(\mathcal{A}) = \frac{1}{2}$ , which is non-negligible.

It remains to show that our distinguisher is PPT. This is the case because it just needs to compute  $\left(\frac{h}{p}\right) = h^{\frac{p-1}{2}}$  for three different elements  $h$  of  $\mathbb{G}$ . Computing  $h^{\frac{p-1}{2}}$  can be done by fast exponentiation, resulting in at most  $\log(p)$  multiplications in  $\mathbb{Z}_p$ . Each such multiplication takes a time polynomial in  $\log(p)$ , and so our distinguisher  $\mathcal{A}$  is indeed polynomial time (in  $\log(p)$ ).

*Remark.* The same reasoning can be adapted if the cardinality of the cyclic group  $\mathbb{G}$  is  $n = km$  for some small  $k$  (and any  $m$ ). In that case, we would have that  $(g^a)^m$  is uniformly distributed among  $\{1, g^m, g^{2m}, \dots, g^{(k-1)m}\}$  if  $a$  is uniform, and computing  $(g^a)^m$  gives us the value of  $a \bmod k$ . We then can check whether  $a \bmod k$  and  $b \bmod k$  are coherent with  $ab \bmod k$ . In the DDH case, this will always be coherent whereas in the uniform case, this will be coherent only with probability  $\frac{k-1}{k}$ . We hence obtain a distinguisher with advantage  $\frac{k-1}{k}$  (which is non negligible for any  $k \geq 2$ ) and whose computation time is  $\Theta(k \cdot \text{poly}(\log(n)))$ . So if  $k$  is polynomial in  $\log(n)$ , this gives us a polynomial time distinguisher with non negligible advantage. This is why, in the next question, we consider of group of prime cardinality.

But this implies knowing a factorisation of  $n$ .

**Exercise 3.**

Attacking the DLG problem

Let  $\mathbb{G}$  be a cyclic group generated by  $g$ , of (known) prime order  $p$ , and let  $h$  be an element of  $\mathbb{G}$ . Let  $F : \mathbb{G} \rightarrow \mathbb{Z}_p$  be a nonzero function, and let us define the function  $H : \mathbb{G} \rightarrow \mathbb{G}$  by  $H(\alpha) = \alpha \cdot h \cdot g^{F(\alpha)}$ . We consider the following algorithm (called *Pollard  $\rho$  Algorithm*).

**Pollard  $\rho$  Algorithm****Input:**  $h, g \in \mathbb{G}$ **Output:**  $x \in \{0, \dots, p-1\}$  such that  $h = g^x$  or FAIL.

1.  $i \leftarrow 1$
2.  $x \leftarrow 0, \alpha \leftarrow h$
3.  $y \leftarrow F(\alpha); \beta \leftarrow H(\alpha)$
4. **while**  $\alpha \neq \beta$  **do**
5.    $x \leftarrow x + F(\alpha) \bmod p; \alpha \leftarrow H(\alpha)$
6.    $y \leftarrow y + F(\beta) \bmod p; \beta \leftarrow H(\beta)$
7.    $y \leftarrow y + F(\beta) \bmod p; \beta \leftarrow H(\beta)$
8.    $i \leftarrow i + 1$
9. **end while**
10. **if**  $i < p$  **then**
11.   **return**  $(x - y)/i \bmod p$
12. **else**
13.   **return** FAIL
14. **end if**

To study this algorithm, we define the sequence  $(\gamma_i)$  by  $\gamma_1 = h$  and  $\gamma_{i+1} = H(\gamma_i)$  for  $i \geq 1$ .

1. Show that in the **while** loop from Steps 4 to 9 of the algorithm, we have  $\alpha = \gamma_i = g^x h^i$  and  $\beta = \gamma_{2i} = g^y h^{2i}$ .

$\text{☞}$  We check these identities by induction on  $i \geq 1$ . For  $i = 1$ , they are satisfied since from lines 1 to 3 of the algorithm, we have  $x = 0, \alpha = h, y = F(h)$ , and  $\beta = H(h) = g^y h^2$ .

Now, let  $i \geq 1$  and denote by  $x_i, \alpha_i, y_i, \beta_i$  the values taken by  $x, \alpha, y, \beta$  at the beginning of the  $i$ -th iteration of the **while** loop. We assume that the identities  $\alpha_i = \gamma_i = g^{x_i} h^i$  and  $\beta_i = \gamma_{2i} = g^{y_i} h^{2i}$  hold.

At the end of the  $i$ -th iteration (or the beginning of the  $i+1$ -th), we have  $x_{i+1} = x_i + F(\alpha_i) \bmod p$ , and  $\alpha_{i+1} = H(\alpha_i) = \alpha_i \cdot h \cdot g^{F(\alpha_i)} = (g^{x_i} \cdot h^i) \cdot h \cdot g^{F(\alpha_i)} = g^{x_i + F(\alpha_i)} \cdot h^{i+1} = g^{x_{i+1}} \cdot h^{i+1}$ . We also have  $\beta_{i+1} = H(H(\beta_i)) = H(\beta_i \cdot h \cdot g^{F(\beta_i)}) = \beta_i \cdot h^2 \cdot g^{F(\beta_i)} \cdot g^{F(H(\beta_i))} = g^{y_i} \cdot g^{F(\beta_i)} \cdot g^{F(H(\beta_i))} \cdot h^{2i+2}$ , and  $y_{i+1} = y_i + F(\beta_i) + F(H(\beta_i))$ , hence the identity  $\beta_{i+1} = g^{y_{i+1}} h^{2i+2}$ .

2. Show that if this loop terminates with  $i < p$ , then the algorithm returns the discrete logarithm of  $h$  in basis  $g$ .

$\text{☞}$  When the loop finishes, we have  $\alpha = \beta$  and according to Question 1, this gives  $g^x h^i = g^y h^{2i}$ , thus  $h^i = g^{x-y}$ . If furthermore the loop finishes with  $i < p$  (note that  $i > 0$ ), then since  $p$  is prime,  $i$  is invertible modulo  $p$  and  $h = g^u$  where  $u = (x - y)/i \bmod p$ .


3. Let  $j$  be the smallest integer such that there exists  $k < j$  such that  $\gamma_j = \gamma_k$ . Show that  $j \leq p + 1$  and that the loop ends with  $i < j$ .

$\text{☞}$  The sequence  $(\gamma_i)$  has its values in the finite group  $\mathbb{G}$  of cardinality  $p$ . By the pigeonhole principle, there exist two indices  $k < j \leq p + 1$  such that  $\gamma_k = \gamma_j$ ; then, since  $(\gamma_i)$  is defined by  $\gamma_{i+1} = H(\gamma_i)$ , this sequence repeats with period a divisor of  $j - k$ .

*Remark:* we have  $\gamma_{k+t} = \gamma_{j+t}$  for any integer  $t \geq 0$ . This leads to representing the values of the sequence in a shape which looks like the letter  $\rho$ , hence the name of the algorithm.

From Question 1, we see that the algorithm simultaneously computes the values of  $\gamma_i$  and  $\gamma_{2i}$  and returns the first index  $i$  for which  $\gamma_i = \gamma_{2i}$ . Since the sequence repeats with period  $j - k$ , considering the smallest multiple  $i$  of  $j - k$  that is greater or equal to  $k$ , namely  $i = (j - k) \lceil \frac{k}{j - k} \rceil$ , we have that  $\gamma_i = \gamma_{2i}$ , since  $i \geq k$  and  $2i - i = i$  is a multiple of the period  $j - k$ . Besides, the sequence  $k, k + 1, \dots, k + (j - k - 1)$  contains a multiple of  $j - k$ , so that  $i \leq j - 1$  (we can also deduce it from the formula above for  $i$ ).

4. Show that if  $F$  is a random function, then the average execution time of the algorithm is in  $O(p^{1/2})$  multiplications in  $\mathbb{G}$ .

 If  $H : \mathbb{G} \rightarrow \mathbb{G}$  is a random function, according to the birthday paradox, the expected number of elements of the sequence  $(\gamma_i)$  needed to obtain two identical values is approximately  $\sqrt{\pi p/2}$ . Since every iteration of the **while** loop uses a constant number of multiplications in  $\mathbb{G}$ , the result follows.